

# NORMAL SAMPLES WITH LINEAR CONSTRAINTS AND GIVEN VARIANCES

D. A. S. FRASER

**1. Summary.** In *Biometrika* (1948) a paper [1] by H. L. Seal contained a theorem applying to “ $n$  random variables normally distributed about zero mean with unit variance, these variables being connected by means of  $k$  linear relations.”<sup>1</sup> Arising from this is the question of how to obtain a set of normal variates connected by  $k$  linear relations and such that each variate has unit variance: or, more generally, connected by  $k$  linear relations and such that each variate has a given variance. The procedure for obtaining such a set of variates when existent from a set of independent normal deviates with unit variances is given in §5. In §§2, 3 and 4, we shall consider various conditions necessary for the existence and construction of such a set.

**2. Normal distribution in a linear subspace.** Let  $(x_1, x_2, \dots, x_n)$  designate a point in an  $n$ -dimensional Euclidean space  $R^n$ . A set of variates  $x_1, x_2, \dots, x_n$  can be considered as a random point in  $R^n$ . In the present problem we shall assume the set has a multivariate normal distribution.

Consider  $k$  homogeneous and independent linear relations

$$\sum_{j=1}^n a_{pj}x_j = 0 \quad (p = n - k + 1, \dots, n).$$

A variate satisfying these relations and these only will belong to an  $(n - k)$ -dimensional linear subspace. By taking linear combinations of the above  $k$  relations, an equivalent set of  $k$  relations can be obtained such that they are orthogonal and normalized:

$$\sum_{j=1}^n b_{pj}x_j = 0 \quad (p = n - k + 1, \dots, n),$$

and

$$\sum_{j=1}^n b_{pj}b_{qj} = \delta_{pq}.$$

By adding  $n - k$  rows, the matrix  $\|b_{pj}\|$  can be completed to an  $n$  by  $n$  matrix  $\|b_{ij}\|$  ( $i, j, = 1, 2, \dots, n$ ) satisfying the orthogonality conditions

$$\sum_{k=1}^n b_{ik}b_{jk} = \delta_{ij}.$$

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Received July 8, 1950.

<sup>1</sup>It is to be noted that the statement of the theorem in [1] is incorrect. The theorem applies to the *residuals* of  $n$  normal variables after fitting  $k$  linear constraints.

This matrix can now be considered as the matrix of an orthogonal rotation of  $n$ -space. Consider coordinates  $y_1, y_2, \dots, y_n$  with respect to the new axes: then

$$y_i = \sum_{j=1}^n b_{ij}x_j.$$

Since normality is invariant under linear transformations, a set of normally distributed  $x$  variates yields a set of normally distributed  $y$  variates, and conversely. A set of variates  $(x_1, x_2, \dots, x_n)$  satisfying  $k$  linear relations

$$\sum_{j=1}^n a_{pj}x_j = 0$$

is transformed by the above to set a of variates  $(y_1, y_2, \dots, y_{n-k})$  satisfying no linear constraints where  $y_{n-k+1}, \dots, y_n$  are identically zero.

**3. Conditions on the variance.** We have solved the problem of linear constraints by working in an  $(n - k)$ -dimensional subspace. How do we interpret in this subspace the original variance conditions:

$$\text{var} \{x_i\} = v_i \quad (i = 1, 2, \dots, n),$$

with

$$y_r = \sum_{j=1}^n b_{rj}x_j,$$

$$x_i = \sum_{r=1}^{n-k} b_{ri}y_r.$$

Each  $x_i$  is seen to be a linear combination of the  $n - k$  variates  $y_r$  and consequently the variance of  $x_i$  can be expressed in terms of the elements of the variance covariance matrix of the  $y_r$ . Consider now a multivariate normal distribution in the subspace with covariance matrix  $\|\tau_{rs}\|$  with respect to the axes  $y_1, y_2, \dots, y_{n-k}$ .

Thus the variance conditions after rotation into the subspace become

$$\sum_{r,s=1}^{n-k} b_{ri}\tau_{rs}b_{si} = v_i \quad (i = 1, 2, \dots, n).$$

**4. Existence.** Our problem has now reduced itself to that of finding a multivariate normal distribution in  $n - k$  dimensions with covariance matrix  $\|\tau_{rs}\|$  such that

$$\sum_{r,s=1}^{n-k} b_{ri}\tau_{rs}b_{si} = v_i \quad (i = 1, 2, \dots, n),$$

or

$$\sum_{r,s=1}^{n-k} c_{ri}\tau_{rs}c_{si} = 1 \quad (i = 1, 2, \dots, n)$$

where  $c_{ri} = v_i^{-\frac{1}{2}}b_{ri}$ .

We have  $n$  equations with  $\binom{n-k+1}{2}$  unknowns. If  $\binom{n-k+1}{2} \geq n$ , a solution will exist and can best be obtained by solving the equation directly. If  $\binom{n-k+1}{2} < n$ , an application of linear regression theory would be indicated.

To find a matrix  $\|\tau_{rs}\|$ , if one exists, is equivalent to finding a generalized ellipsoid

$$\sum_{r,s=1}^{n-k} z_r \tau_{rs} z_s = 1$$

passing through the  $n$  points

$$(c_{1i}, \dots, c_{n-k,i}) \quad (i = 1, 2, \dots, n).$$

This is accomplished using linear regression theory by fitting to the constant 1 the functions  $z_r z_s$  ( $r, s = 1, 2, \dots, n-k$ ) for the  $n$  "sample" values given above of the vector  $(z_1, z_2, \dots, z_{n-k})$ . If the sum of squares for residuals is zero then a quadratic surface exists. However, to have a solution to our distribution problem, the matrix of the quadratic form must be positive. If it is not positive definite, then our variance conditions have imposed a further linear constraint on the set of variates.

**5. Conclusions.** The problem may be stated: to find normal variables  $x_1, x_2, \dots, x_n$  satisfying  $k$  homogeneous and independent linear relations

$$\sum_{j=1}^n a_{pj} x_j = 0 \quad (p = n - k + 1, \dots, n),$$

and with

$$\text{var} \{x_i\} = v_i \quad (i = 1, 2, \dots, n).$$

The solution can be described in five steps.

**5.1.** Find a matrix  $\|b_{pj}\|$  with  $p = n - k + 1, \dots, n$  and  $j = 1, 2, \dots, n$  with orthogonal and normalized rows equivalent to  $\|a_{pj}\|$  as described in §2.

**5.2.** Complete  $\|b_{pj}\|$  to an orthogonal matrix  $\|b_{ij}\|$  ( $i, j = 1, 2, \dots, n$ ).

**5.3.** Find a quadratic equation

$$\sum_{r,s=1}^{n-k} z_r \tau_{rs} z_s = 1$$

satisfied by the  $n$  points

$$(b_{1i} v_i^{-\frac{1}{2}}, \dots, b_{n-k,i} v_i^{-\frac{1}{2}}) \quad (i = 1, 2, \dots, n),$$

if it exists, directly or by regression theory as in §4. If the equation does not exist or if it exists with a non-positive matrix then the problem has no solution.

**5.4.** If the matrix  $\|\tau_{rs}\|$  is positive, then find random variates  $y_1, y_2, \dots, y_{n-k}$  with zero means and  $\|\tau_{rs}\|$  as covariance matrix. (If  $\|\tau_{rs}\|$  is positive

definite, take the square root matrix of  $||\tau_{rs}||$  and apply as a linear transformation to  $n - k$  independent normal variates with means 0 and variances 1. If positive but not definite, then the previous method will work in a subspace of  $y_1, y_2, \dots, y_{n-k}$ .)

5.5. Obtain the set of  $x$  variates, thus solving the problem, by applying the transformation

$$x_i = \sum_{r=1}^{n-k} b_{ri} y_r$$

to the  $y$  variates obtained in 5.4.

#### REFERENCES

1. H. L. Seal, *A note on the  $\chi^2$  smooth test*, *Biometrika*, vol. 35 (1948), 202.

*University of Toronto*